

Seemingly unrelated Regressions (SUR)

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Overview

- What are SUR-models?
- 1-Way Model
- 2-Way Model
- Example
- Simultaneous Equations

SUR in general

- Concept by Arnold Zellner from 1962
- Basic idea: Error terms of different equations are correlated amongst each other
 - *f.ex.: demand of different firms*
or for different products (substitutes)
- Same number of explanatory variables in each equation, not necessarily same X

SUR with panel-data

- *“Same same but different!”*
- Feasible GLS is used to estimate the variance-covariance matrix and parameter estimates
- The process is iterated until the errors are minimized
- Variables could be transformed with a form of Cochrane-Orchutt correction to model autocorrelation

1-Way model

A set of equations ...

$$y_1 = Z_1 \delta_1 + u_1$$

\vdots

$$y_M = Z_M \delta_M + u_M$$

$$y_j = Z_j \delta_j + u_j$$

1-Way model

... with error terms ...

$$u_j = Z_{\mu} \mu_j + v_j$$

where:

$$Z_{\mu} = (I_N \otimes I_T)$$

$$\mu_j = (\mu_{1j}, \mu_{2j}, \dots, \mu_{Nj})$$

$$v_j = (v_{11j}, \dots, v_{1Tj}, \dots, v_{N1j}, \dots, v_{NTj})$$

1-Way model

... has the following covariance matrices:

$$\begin{aligned}\Omega_{jl} &= E(u_j u_l') \\ &= \sigma_{\mu_{jl}}^2 (I_N \otimes J_T) + \sigma_{\nu_{jl}}^2 (I_N \otimes I_T)\end{aligned}$$

Variance-Covariance Matrix:

$$\begin{aligned}\Omega &= E(uu') \\ &= \sum_{\mu} \otimes (I_N \otimes J_T) + \sum_{\nu} \otimes (I_N \otimes I_T)\end{aligned}$$

1-Way model

$$\begin{aligned}
 &= \begin{pmatrix} \sigma_{\mu_{11}}^2 & \dots & \dots & \sigma_{\mu_{1j}}^2 \\ \sigma_{\mu_{21}}^2 & \ddots & & \\ \vdots & & \ddots & \\ \sigma_{\mu_{j1}}^2 & & & \sigma_{\mu_{jl}}^2 \end{pmatrix} \otimes \begin{pmatrix} 1 & \dots & 1 & & & \\ \vdots & & \vdots & & 0 & \\ 1 & \dots & 1 & & & \\ & & & 1 & \dots & 1 \\ & & & \vdots & & \vdots \\ & & & 1 & \dots & 1 \end{pmatrix} \\
 &+ \begin{pmatrix} \sigma_{v_{11}}^2 & \dots & \dots & \sigma_{v_{1j}}^2 \\ \sigma_{v_{21}}^2 & \ddots & & \\ \vdots & & \ddots & \\ \sigma_{v_{j1}}^2 & & & \sigma_{v_{jl}}^2 \end{pmatrix} \otimes \begin{pmatrix} 1 & & & & & \\ & \ddots & & & & \\ & & \ddots & & & \\ & & & \ddots & & \\ & & & & \ddots & \\ & & & & & 1 \end{pmatrix}
 \end{aligned}$$

1-Way model

now we replace $J_T = T\bar{J}_T$ and $I_T = E_T + \bar{J}_T$

$$\begin{aligned}\Omega &= T \sum_{\mu} \otimes (I_N \otimes \bar{J}_T) + \sum_{\nu} \otimes (I_N \otimes (E_T + \bar{J}_T)) \\ &= (T \sum_{\mu} + \sum_{\nu}) \otimes (I_N \otimes \bar{J}_T) + \sum_{\nu} \otimes (I_N \otimes E_T) \\ &= \sum_1 \otimes P + \sum_{\nu} \otimes Q, \quad \text{with} \quad Q = I_{NT} - P\end{aligned}$$

in order to be able to invert Ω

$$\begin{aligned}\Omega^r &= \sum_1^r \otimes P + \sum_{\nu}^r \otimes Q \\ \Omega^{-1} &= \sum_1^{-1} \otimes P + \sum_{\nu}^{-1} \otimes Q \\ \Omega^{-1/2} &= \sum_1^{-1/2} \otimes P + \sum_{\nu}^{-1/2} \otimes Q\end{aligned}$$

1-Way model

VC matrix can be estimated using OLS or within-type residuals

$$\hat{\Sigma}_1 = \frac{U'PU}{N}, \quad \hat{\Sigma}_\mu = \frac{U'QU}{N(T-1)}$$

if $N, T \rightarrow \infty$, the Within estimator is asymptotically efficient and has the same asymptotic VC matrix as the GLS estimator

2-Way model

Now we include a time effect:

$$u_j = Z_\mu \mu_j + Z_\lambda \lambda_j + v_j$$

where

$$Z_\mu = (I_N \otimes l_T)$$

$$Z_\lambda = (l_T \otimes I_N)$$

$$\mu_j = (\mu_{1j}, \mu_{2j}, \dots, \mu_{Nj})$$

$$\lambda_j = (\lambda_{1j}, \lambda_{2j}, \dots, \lambda_{Tj})$$

$$v_j = (v_{11j}, \dots, v_{1Tj}, \dots, v_{N1j}, \dots, v_{NTj})$$

2-Way model

The error term matrix has then the following form:

$$\Omega = E(uu') = \Sigma_{\mu} \otimes (I_N \otimes J_T) + \Sigma_{\lambda} \otimes (J_N \otimes I_T) + \Sigma_{\nu} \otimes (I_N \otimes I_T)$$

Ω has to be decomposed to invert:

$$\begin{aligned}\Omega &= T \Sigma_{\mu} \otimes (I_N \otimes \bar{J}_T) + N \Sigma_{\lambda} \otimes (\bar{J}_N \otimes I_T) + \Sigma_{\nu} \otimes ((E_N + \bar{J}_N) \otimes (E_T + \bar{J}_T)) \\ &= (T \Sigma_{\mu} + \Sigma_{\nu}) \otimes (I_N \otimes \bar{J}_T) + (N \Sigma_{\lambda} + \Sigma_{\nu}) \otimes (\bar{J}_N \otimes I_T) + \Sigma_{\nu} \otimes (I_N \otimes E_T) \\ &\Rightarrow \sum_{i=1}^4 \Lambda_i \otimes Q_i\end{aligned}$$

$$\Lambda_1 = \Sigma_{\nu}, \quad \Lambda_2 = T \Sigma_{\mu} + \Sigma_{\nu}, \quad \Lambda_3 = N \Sigma_{\lambda} + \Sigma_{\nu}, \quad \Lambda_4 = T \Sigma_{\mu} + N \Sigma_{\lambda} + \Sigma_{\nu}$$

2-Way model

Having done this, we can introduce any arbitrary exponent r

$$\begin{aligned}\Omega^r &= \sum_{i=1}^4 \Lambda_i^r \otimes Q_i \\ \Rightarrow \sum_{i=1}^4 \Lambda_i^{-1} \otimes Q_i, & \quad or \quad \Rightarrow \sum_{i=1}^4 \Lambda_i^{-1/2} \otimes Q_i\end{aligned}$$

individual components of Ω can be estimated

$$\hat{\Sigma}_v = \frac{U' Q_1 U}{(N-1)(T-1)}, \quad \hat{\Lambda}_2 = \frac{U' Q_2 U}{(N-1)}, \quad \hat{\Lambda}_3 = \frac{U' Q_3 U}{(T-1)}$$

with U derived from OLS or within estimation

Example

Stata commands

- *xtsur* → 1-way random effects estimation
- Alternatives:
 - sureg* + dummies for individual and/or time effects + constraints ($\sum \mu_i = 0, \sum \lambda_t = 0$)
 - *noconstant*-option does not work with *sureg*

Source: FAOSTAT (<http://faostat.fao.org/>)

Example: 1-Way estimation

•	variable		egg1		milk1
•	-----+-----				
•	1chick		.41690575		
•			0.0000		
•	1chick_meat		.38514723		
•			0.0000		
•	1gdp		-.02212248		.1390674
•			0.4905		0.0000
•	1cat				.54145784
•					0.0000
•	1cat_meat				.32742093
•					0.0000
•	_cons		5.4959816		.66752561
•			0.0000		0.3943
•	-----+-----				
•	N		354		354
•	theta		.94029847		.96263607
•	-----				
•					Legend: b/p

Example: 1-Way estimation

```

• . xtsur (llegg lchick lchick_meat lgdp) ///
• >      (lmilk lcat lcat_meat lgdp)
•
• Seemingly unrelated regression (SUR) in unbalanced panel data set
• One-way random effect estimation:
• -----
• Number of Group variable:      1                Number of obs      =      354
• Panel variable: code           Number of eqn      =      2
• Time variable : year          Number of groups   =      1
• -----
•
•      |      Coef.   Std. Err.      z    P>|z|      [95% Conf. Interval]
• -----+-----
• llegg |
•   lchick |   .7816699   .0517333   15.11  0.000   .6802745   .8830653
•   lchick_meat |   .3962238   .0507861    7.80  0.000   .2966848   .4957627
•   lgdp |   .1567003   .0329094    4.76  0.000   .0921991   .2212014
• -----+-----
• lmilk |
•   lcat |   .5815703   .0433592   13.41  0.000   .4965877   .6665528
•   lcat_meat |   .312817    .0528716    5.92  0.000   .2091906   .4164434
•   lgdp |   .1613396   .0306559    5.26  0.000   .1012552   .2214241
• -----

```


Example: 2-Way estimation

•	variable		egg2		milk2
•					
•	1chick		.42192118		
•			0.0000		
•	1chick_meat		.37935621		
•			0.0000		
•	1gdp		.00337928		.34067258
•			0.9396		0.0000
•	t2		-.00512125		-.04554634
•			0.6833		0.0003
•	t3		-.013413		-.08610724
•			0.4128		0.0000
•	1cat				.59157518
•					0.0000
•	1cat_meat				.3033061
•					0.0000
•	_cons		5.3030452		-1.4259562
•			0.0000		0.1074
•					
•	N		354		354
•	theta		.94067432		.96258285
•					
•					Legend: b/p

Example: 2-Way estimation

```
• . xtsur (llegg lchick lchick_meat lgdp t*) ///  
• > (lmilk lcat lcat_meat lgdp t*)  
•  
• One-way random effect estimation:  
• -----  
• | Coef. Std. Err. z P>|z| [95% Conf. Interval]  
• -----+-----  
• llegg |  
• lchick | .426133 .0479355 8.89 0.000 .3321812 .5200847  
• lchick_meat | .3713545 .0403583 9.20 0.000 .2922537 .4504552  
• lgdp | -.0648342 .0407461 -1.59 0.112 -.1446952 .0150268  
• t1 | 5.912767 .4476631 13.21 0.000 5.035363 6.79017  
• t2 | 5.915545 .4509307 13.12 0.000 5.031737 6.799353  
• t3 | 5.918376 .4555626 12.99 0.000 5.02549 6.811262  
• -----+-----  
• lmilk |  
• lcat | .4845353 .0602053 8.05 0.000 .3665351 .6025355  
• lcat_meat | .2245834 .0509356 4.41 0.000 .1247515 .3244153  
• lgdp | .2701145 .0545115 4.96 0.000 .1632738 .3769551  
• t1 | 1.585068 .8622317 1.84 0.066 -.1048753 3.275011  
• t2 | 1.548979 .8661407 1.79 0.074 -.1486259 3.246583  
• t3 | 1.522254 .871834 1.75 0.081 -.1865097 3.231017  
• -----+-----
```

Simultaneous Equations

- Differs from SUR, as there are rhs endogenous variables in the system of equations

$$Z = \text{diag}[Z_j] \Rightarrow Z_j = [Y_j, X_j]$$

$$\Rightarrow \begin{pmatrix} Z_1 & & & \\ & Z_2 & & \\ & & \ddots & \\ & & & Z_M \end{pmatrix} = \begin{pmatrix} [Y_1, X_1] & & & \\ & \ddots & & \\ & & \ddots & \\ & & & [Y_M, X_M] \end{pmatrix}$$

- The VC matrix has the same form as in SUR!

Simultaneous Equations

System of equations:

$$\tilde{y} = \tilde{Z}\delta + \tilde{u}$$

$$\bar{y} = \bar{Z}\delta + \bar{u}$$

where

$$\tilde{y} = (I_M \otimes Q)y$$

$$\bar{y} = (I_M \otimes P)y$$

$$\tilde{Z} = (I_M \otimes Q)Z$$

$$\bar{Z} = (I_M \otimes P)Z$$

$$\tilde{u} = (I_M \otimes Q)u$$

$$\bar{u} = (I_M \otimes P)u$$

$$Q = I_{NT} - P,$$

$$P = I_N \otimes \bar{J}_T$$

Simultaneous Equations

- Within 3SLS-estimator

$$\tilde{y} = \tilde{Z}\delta + \tilde{u} \Rightarrow \tilde{\delta}_{W3SLS}$$

- Between 3SLS-estimator

$$\bar{y} = \bar{Z}\delta + \bar{u} \Rightarrow \hat{\delta}_{B3SLS}$$

- (Random) Error component 3SLS-estimator

$$\hat{\delta}_{EC3SLS} = W_1 \tilde{\delta}_{W3SLS} + W_2 \hat{\delta}_{B3SLS}$$