

Exercise: Show that the set \mathbb{C}^n of *n*-dimensional complex vectors together with vector addition defined by

$$z+w = \begin{pmatrix} z_1 \\ \vdots \\ z_n \end{pmatrix} + \begin{pmatrix} w_1 \\ \vdots \\ w_n \end{pmatrix} = \begin{pmatrix} z_1 + w_1 \\ \vdots \\ z_n + w_n \end{pmatrix}$$

is a commutative group.

<u>Exercise</u>: Show that the set \mathbb{C}^n together with vector addition defined as above and scalar multiplication defined by

$$\lambda z = \lambda \begin{pmatrix} z_1 \\ \vdots \\ z_n \end{pmatrix} = \begin{pmatrix} \lambda z_1 \\ \vdots \\ \lambda z_n \end{pmatrix}$$

is a complex vector space, i.e.,

$$\lambda(\mu z) = (\lambda \mu)z,$$

$$1z = z,$$

$$\lambda(z+w) = (\lambda z) + (\lambda w),$$

$$(\lambda + \mu)z = (\lambda z) + (\mu z).$$
 FV

Exercise: Show that the inner product of two complex vectors z and w defined by

$$\langle z, w \rangle = w * z = \overline{(w_1, \dots, w_n)} \begin{pmatrix} z_1 \\ \vdots \\ z_n \end{pmatrix} = \sum_{t=1}^n z_t \overline{w_t},$$

satisfies

FG

(i)
$$\langle v + w, z \rangle = \langle v, z \rangle + \langle w, z \rangle$$
,
(ii) $\langle z, v + w \rangle = \langle z, v \rangle + \langle z, w \rangle$,
(iii) $\langle \lambda z, w \rangle = \lambda \langle z, w \rangle$,
(iv) $\langle v, \lambda w \rangle = \overline{\lambda} \langle v, w \rangle$,
(v) $\langle z, w \rangle = \overline{\langle w, z \rangle}$,
(vi) $\langle z, z \rangle \ge 0$,
(vii) $\langle z, z \rangle = 0 \Leftrightarrow z = 0$.

Two complex vectors z and w are said to be orthogonal, if $\langle z, w \rangle = 0.$

Exercise: Show that the norm of a complex vector z defined by

$$\left\|z\right\| = \sqrt{\left\langle z, z\right\rangle} = \sqrt{\sum_{t=1}^{n} z_t \overline{z_t}} = \sqrt{\sum_{t=1}^{n} \left|z_t\right|^2}$$

satisfies

(i)
$$\|\lambda z\| = |\lambda| \|z\|$$
,
(ii) $\|z + w\| \le \|z\| + \|w\|$. FN

Hint: Use the Cauchy-Schwarz Inequality $|\langle z, w \rangle| \le ||z|| ||w||.$

Exercise: Prove the Pythagorean theorem

$$|z,w\rangle = 0 \implies ||z+w||^2 = ||z||^2 + ||w||^2.$$
 FP

Sinusoid: $g(t) = R \sin(\omega t + \phi)$ Parameters: R (amplitude) ω (frequency) ϕ (phase)

A sinusoid is periodic with period $p = \frac{2\pi}{\omega}$ because $g(t + \frac{2\pi}{\omega}) = R \sin(\omega(t + \frac{2\pi}{\omega}) + \phi)$

$$g(t + \frac{2\pi}{\omega}) = R \sin(\omega (t + \frac{2\pi}{\omega}) + \phi)$$

= $R \sin(\omega t + \phi + 2\pi)$
= $R \sin(\omega t + \phi)$
= $g(t)$.

For fixed *n*, the frequencies $\omega_k = \frac{2\pi}{n} \cdot k$, $k=0,...,[\frac{n}{2}]$ are called Fourier frequencies.

$$\omega_1 = \frac{2\pi}{n} \cdot 1$$
 implies a period of $p_1 = \frac{2\pi}{\omega_1} = \frac{2\pi}{\frac{2\pi}{n-1}} = n$.
 $\omega_2 = \frac{2\pi}{n} \cdot 2$ implies a period of $p_2 = \frac{2\pi}{\omega_2} = \frac{2\pi}{\frac{2\pi}{n-2}} = \frac{n}{2}$.

Exercise: Use the Euler relation

$$e^{i\omega} = \cos(\omega) + i\sin(\omega)$$

to show that $e^{i(\omega+2\pi)} = e^{i\omega}$, $\overline{e^{i\omega}} = e^{-i\omega}$, $e^{i(\omega_j - \omega_k)n} = 1$.

(in 1)

The *n* vectors

$$e_k = \frac{1}{\sqrt{n}} \begin{pmatrix} e^{i\omega_k 1} \\ \vdots \\ e^{i\omega_k n} \end{pmatrix}, k = 0, \dots n-1,$$

constitute an orthonormal basis for \mathbf{C}^n because

$$\begin{split} \left\langle e_{k}, e_{k} \right\rangle &= \frac{1}{n} \sum_{t=1}^{n} e^{i\omega_{k}t} \overline{e^{i\omega_{k}t}} = \frac{1}{n} \sum_{t=1}^{n} e^{i\omega_{k}t} e^{-i\omega_{k}t} = \frac{1}{n} \sum_{t=1}^{n} e^{0} = 1, \\ \left\langle e_{j}, e_{k} \right\rangle &= \frac{1}{n} \sum_{t=1}^{n} e^{i\omega_{j}t} \overline{e^{i\omega_{k}t}} = \frac{1}{n} \sum_{t=1}^{n} e^{i\omega_{j}t} e^{-i\omega_{k}t} = \frac{1}{n} \sum_{t=1}^{n} (e^{i(\omega_{j}-\omega_{k})})^{t} \\ &= \frac{1}{n} e^{i(\omega_{j}-\omega_{k})} \sum_{t=0}^{n-1} (e^{i(\omega_{j}-\omega_{k})})^{t} = \frac{1}{n} e^{i(\omega_{j}-\omega_{k})} \frac{1-e^{i(\omega_{j}-\omega_{k})n}}{1-e^{i(\omega_{j}-\omega_{k})}} \\ &= 0 \quad \text{if } j \neq k. \end{split}$$

Thus, any $x \in \mathbb{C}^n$ has a representation of the form

FE

FB

$$x = \sum_{k=0}^{n-1} \lambda_k e_k \; .$$

Taking inner products of each side we obtain

$$\langle x, e_k \rangle = \left\langle \sum_{j=0}^{n-1} \lambda_j e_j, e_k \right\rangle = \sum_{j=0}^{n-1} \lambda_j \left\langle e_j, e_k \right\rangle = \lambda_k \left\langle e_k, e_k \right\rangle = \lambda_k,$$

$$\lambda_k = \left\langle x, e_k \right\rangle = e_k^* x = \frac{1}{\sqrt{n}} \sum_{t=1}^n x_t e^{-i\omega_k t} ,$$

$$\|x\|^2 = \left\langle x, x \right\rangle = \sum_{k=0}^{n-1} \sum_{j=0}^{n-1} \lambda_k \overline{\lambda}_j \left\langle e_k, e_j \right\rangle = \sum_{k=0}^{n-1} \lambda_k \overline{\lambda}_k = \sum_{k=0}^{n-1} |\lambda_k|^2 .$$

Exercise: Show that
$$\frac{1}{n} \sum_{t=1}^{n} (x_t - \overline{x})^2 = \frac{1}{n} \sum_{k \neq 0} |\lambda_k|^2$$
. FF

Note: For a given time series $x_1, ..., x_n$ observed at times 1, ..., n or time intervals (0,1), ..., (n-1,n), the size of $|\lambda_k|$ therefore indicates how much of the sample variance can be attributed to frequency ω_k .

FS

Suppose that
$$x \in \mathbb{R}^n$$
. If $1 \le k < \frac{n}{2}$, then
 $e^{i\omega_{n-k}t} = e^{i\frac{2\pi(n-k)}{n}t} = e^{i(2\pi - \frac{2\pi k}{n})t} = e^{-i\frac{2\pi k}{n}t} = e^{-i\omega_k t} = \overline{e^{i\omega_k t}}$,
 $\lambda_{n-k} = \frac{1}{\sqrt{n}} \sum_{t=1}^n x_t e^{-i\omega_{n-k}t} = \frac{1}{\sqrt{n}} \sum_{t=1}^n x_t \overline{e^{-i\omega_k t}} = \frac{1}{\sqrt{n}} \sum_{t=1}^n x_t e^{-i\omega_k t} = \overline{\lambda_k}$,
 $\lambda_k \frac{e^{i\omega_k t}}{\sqrt{n}} + \lambda_{n-k} \frac{e^{i\omega_{n-k}t}}{\sqrt{n}} = \frac{\lambda_k}{\sqrt{n}} e^{i\omega_k t} + \frac{\overline{\lambda_k}}{\sqrt{n}} \overline{e^{i\omega_k t}}$
 $= (a_k + ib_k)(\cos(\omega_k t) + i\sin(\omega_k t))$
 $+ (a_k - ib_k)(\cos(\omega_k t) - i\sin(\omega_k t))$
 $= 2a_k \cos(\omega_k t) - 2b_k \sin(\omega_k t)$
 $= R_k \sin(\phi_k) \cos(\omega_k t) + R_k \cos(\phi_k) \sin(\omega_k t)$
 $= R_k \sin(\omega_k t + \phi_k)$,²

where R_k and ϕ_k are the polar coordinates of $-2b_k+2a_ki$, i.e., $2a_k = R_k \sin(\phi_k), -2b_k = R_k \cos(\phi_k).$

If
$$k = \frac{n}{2}$$
, then
 $\omega_k = \frac{2\pi k}{n} = \pi$, $e^{i\omega_k t} = e^{i\pi t} = \cos(\pi t) + \underbrace{\sin(\pi t)}_{=0} = \sin(\pi t + \frac{\pi}{2})$,
 $\lambda_k = \frac{1}{\sqrt{n}} \sum_{t=1}^n x_t \cos(\pi t) = \frac{1}{\sqrt{n}} \sum_{t=1}^n x_t (-1)^t \in \mathbf{R}$,
 $\frac{\lambda_k}{\sqrt{n}} e^{i\omega_k t} = R_k \cos(\pi t) = R_k \sin(\pi t + \frac{\pi}{2}) = R_k \sin(\omega_k t + \phi_k)$.
If $k = 0$, then
 $\omega_k = \frac{2\pi 0}{n} = 0$, $e^{i\omega_k t} = e^{i0t} = 1$,

If k=0, then

$$\omega_{k} = \frac{2\pi 0}{n} = 0, \ e^{i\omega_{k}t} = e^{i0t} = 1,$$

$$\lambda_{k} = \frac{1}{\sqrt{n}} \sum_{t=1}^{n} x_{t} \cdot 1 = \sqrt{n} \frac{1}{n} \sum_{t=1}^{n} x_{t} = \sqrt{n} \overline{x} \in \mathbf{R},$$

$$\frac{\lambda_{k}}{\sqrt{n}} e^{i\omega_{k}t} = \overline{x}.$$

Thus,

$$x_t = \frac{1}{\sqrt{n}} \left(\lambda_0 + \sum_{k=1}^{n-1} \lambda_k e^{i\omega_k t} \right) = \overline{x} + \sum_{k=1}^{[n/2]} R_k \sin(\omega_k t + \phi_k).$$

² $\sin(\alpha + \beta) = \sin(\alpha)\cos(\beta) + \cos(\alpha)\sin(\beta)$

The **periodogram** of x_1, \ldots, x_n is defined by

$$I(\omega) = \frac{1}{2\pi n} \left| \sum_{t=1}^{n} x_t e^{-i\omega t} \right|^2$$

For $1 \le k < \frac{n}{2}$,

$$I(\omega_k) = \frac{1}{2\pi} |\lambda_k|^2 = \frac{1}{2\pi} (a_k^2 + b_k^2) = \frac{1}{8\pi} ((2a_k)^2 + (-2b_k)^2) = \frac{1}{8\pi} R_k^2.$$

It will be shown later that for any nonzero Fourier frequency ω_k , $I(\omega_k)$ can be written as

$$I(\omega_{k}) = \frac{1}{2\pi} \sum_{j=-(n-1)}^{n-1} \hat{\gamma}(j) e^{-i\omega_{k} j},$$

where

$$\hat{\gamma}(j) = \frac{1}{n} \sum_{t=1}^{n-|j|} (x_t - \bar{x}) (x_{t+|j|} - \bar{x})$$

is the sample autocovariance at lag j.

If the observations $x_1, ..., x_n$ come from a stationary process x, the periodogram $I(\omega)$ may be regarded as a sample analogue of the function

$$f(\omega) = \frac{1}{2\pi} \sum_{j=-\infty}^{\infty} \gamma(j) e^{-i\omega j}$$

which is called the **spectral density** of the process *x*.

The stationarity of the process x implies that all x_t have the same mean and the same variance and the autocovariances

$$\gamma(j) = \operatorname{Cov}(x_t, x_{t-j})$$

depend only on *j* but not on *t*.

Exercise: Show that
$$\int_{-\pi}^{\pi} e^{i\omega k} d\omega = 0$$
 if $k \neq 0$. A0

Assuming that the interchange of summation and integration is justified we can derive the **spectral representation of the autocovariance function** γ of a stationary process x with spectral density f as follows:

$$\int_{-\pi}^{\pi} e^{i\omega k} f(\omega) d\omega = \int_{-\pi}^{\pi} e^{i\omega k} \frac{1}{2\pi} \sum_{j=-\infty}^{\infty} \gamma(j) e^{-i\omega j} d\omega$$
$$= \frac{1}{2\pi} \sum_{j=-\infty}^{\infty} \gamma(j) \int_{-\pi}^{\pi} e^{i\omega(k-j)} d\omega$$
$$= \frac{1}{2\pi} \gamma(k) \int_{-\pi}^{\pi} e^{i\omega(k-k)} d\omega$$
$$= \gamma(k)$$

Remark: Let

$$f_J(\omega) = \sum_{j=-J}^{J} \gamma(j) \cos(\omega j).$$

If

$$g(\omega) = \sum_{j=-\infty}^{\infty} |\gamma(j)| < \infty,$$

then

$$\int_{-\pi}^{\pi} g(\omega) d\omega < \infty$$

and, by the dominant convergence theorem,

$$\lim_{J\to\infty}\int_{-\pi}^{\pi}f_{J}(\omega)d\omega=\int_{-\pi}^{\pi}\lim_{J\to\infty}f_{J}(\omega)d\omega\,,$$

because

$$|f_J(\omega)| \leq \sum_{j=-J}^{J} |\gamma(j) \cos(\omega j)| \leq \sum_{j=-J}^{J} |\gamma(j)| \leq g(\omega).$$

AR

Remark: It follows from

$$f(\boldsymbol{\omega}) = \frac{1}{2\pi} \sum_{j=-\infty}^{\infty} \gamma(j) e^{-i\omega j}$$

and

 $\gamma(j) = \int_{-\pi}^{\pi} e^{i\omega j} f(\omega) d\omega$

that the spectral density f and the autocovariance function γ contain the identical information.

Given *n* observations x_1, \ldots, x_n , the periodogram

$$I(\omega) = \frac{1}{2\pi} \sum_{j=-(n-1)}^{n-1} \hat{\gamma}(j) e^{-i\omega}.$$

is a very erratic estimator for the spectral density, because the sample autocovariances

$$\hat{\gamma}(j) = \frac{1}{n} \sum_{t=1}^{n-|j|} (x_t - \bar{x})(x_{t+|j|} - \bar{x})$$

contain very few products $(x_t - \overline{x})(x_{t+|j|} - \overline{x})$ if |j| is large.

An obvious improvement is to give less weight to the more variable sample autocovariances.

An estimator of the type

$$\hat{f}(\omega) = \frac{1}{2\pi} \sum_{j=-(n-1)}^{n-1} w_j \hat{\gamma}(j) e^{-i\omega j}$$

is called a **weighted covariance estimator**. A widely used estimator is the **Bartlett estimator** which uses the triangle weights

$$w_{j} = \begin{cases} 1 - \frac{|j|}{M} & \text{if } |j| < M, \\ 0 & \text{else.} \end{cases}$$

The truncation point M is an important parameter for controlling the smoothness of the estimator.

An alternative method of smoothing the periodogram is to take weighted averages over neighboring frequencies. A widely used **smoothed periodogram estimator** is the **modified Daniell smoother**, which differs from a simple moving average of the periodogram only in that the first and the last weight are only half as large as the others. Exercise: Spectral analysis of the postwar US GDP

• Create a working directory, say C:\GDPq, for the analysis of the quarterly US GDP.

• Download the real Gross Domestic Product (quarterly, seasonally adjusted) as a csv file from the website of the Federal Reserve Bank of St. Louis into your working directory. The downloaded file **GDPC1.csv** consists of two columns (dates and GDP values).

• Import the data into R and plot the GDP, the log GDP, the differenced log GDP, and the periodogram of the differences.

setwd("C:/GDPq") # comment: set working directory D <- read.csv("GDPC1.csv") # import data d <- D[,1] # 1st column of D: dates v <-D[,2] # 2nd column of D: GDP values d <- as.Date(d) # convert character strings to dates N <- length(v) # N = no. of quarters = length of vector v par(mfrow=c(2,2)) # subsequent plots in 2x2 array par(mar=c(2,2,1,1)) # set narrow margins for plots plot(d,v,pch=20) # plot GDP values against dates y <- log(v); plot(d,y,pch=20) # plot character: solid circle r <- y[2:N]-y[1:(N-1)]; n <- N-1 # n = no. of differences plot(d[2:N],r,pch=20,type="o") # overplot points&lines h <- spec.pgram(r,taper=0,detrend=F,fast=F,plot=F) c <- 2*pi; f <- c*h\$freq # Fourier fr. between 0 and pi pg <- h\$spec/c; plot(f,pg,type="o",pch=20) # periodogr.



• Smooth the periodogram with the modified Daniell smoother.



The higher the span (the total number of terms in the moving average), the smoother the estimate.