

# Estimation and Hypothesis Testing with Unfiltered and Filtered Data

## Ghysels & Osborn Chapter 5

Nora Prean

Econometrics of seasonality, May 2012

# Adjusted versus unadjusted data

- Seasonally adjusted data
  - e.g. structural models designed to capture business cycle fluctuations
- Unadjusted data
  - models should reflect decision-making process of representative agents **facing seasonal fluctuations**
- Question will remain unsolved; BUT researchers should be aware of econometric consequences of using seasonally unadjusted and seasonally adjusted data

# Outlook

1. Misspecification of stochastic structure of seasonality in linear regression models
2. Consequences of using seasonally adjusted (filtered) data
3. Conclusion

# Sources of misspecification

Misspecifying seasonality as deterministic when in fact it is stochastic and nonstationary and vice versa

Assume

$$y_t = \sum_{s_1}^S \delta_{st} \gamma_s + z_t \quad (5.1)$$

where  $\phi(L)z_t = \theta(L)\epsilon_t$ ,  $\epsilon_t \sim \text{iid}(0, \sigma^2)$

- Inappropriate seasonal differencing:

- if seasonality is deterministic (all roots outside unit circle):  
run regression in (5.1)

- corresponding seasonal difference model:

$$\Delta_S y_t = \Delta_S z_t \quad (5.2)$$

- seasonal differences operator  $\Delta_S$  removes seasonal means  
(no longer identified)
- furthermore, the  $\Delta_S$  filter results in  $\Delta_S z_t$  ("overdifferencing"  
of the data, see Maravall, 1995)

## Misspecification, cont.

- Inappropriate seasonal dummies [ $\phi(L)$  has roots on unit circle at seasonal frequencies]:
  - "spurious deterministic seasonality" (Abeyasinghe, 1991, 1994; Franses, Hylleberg, Lee (1995)
  - R2 misleading (estimation of (5.1) might lead to high R2 even if DGP is nonstationary seasonal process)
  - conventional F test for the null  $\gamma_1 = \dots = \gamma_s$  not valid

## Excursus: filtering

- Idea: filter raw data such that seasonal fluctuations disappear from the series
- Long tradition of decomposing time series in mutually orthogonal unobserved components
  - any seasonal adjustment procedure rests on specific decomposition of a series into a **"trend cycle"**, and **seasonal** and **irregular components**
  - multiplicative, pseudo-additive, (log) additive decomposition
- Linear X-11 filter (U.S. Census Bureau, N.B.E.R), for monthly and quarterly data
  - 3 stage procedure, each with several sub-stages, using different MA filters, seasonal moving average filters,...

# Regression models and Filtering

Consider:

$$\begin{aligned}y_t &= \sum_{j=1}^p \alpha_j y_{t-j} + x_t' \beta + \delta_t + \epsilon \\t &= -l, \dots, 0, 1, \dots, T+k\end{aligned}\quad (5.6)$$

regressors non-stochastic, error process  $\epsilon' = (\epsilon_{-l}, \dots, \epsilon_{T+k})$  has known covariance matrix  $E(\epsilon\epsilon') = \Omega$

Assume that data is filtered with a known and linear filter; filter weights  $(v_{-l}, \dots, v_k)$



## Filter matrix $F$

$(T + 1) \times (T + k + l + 1)$  filter matrix  $F$ :

$$F = \begin{pmatrix} v_{-l} & \dots & \dots & v_k & 0 \\ & \ddots & & & \ddots \\ & & & \ddots & \ddots \\ 0 & & v_{-l} & & v_k \end{pmatrix}$$

- Filter matrix  $F$  transforms sample of size  $\mathcal{T} = T + k + l + 1$  into filtered data set with  $T + 1$  observations (at each end of sample data are discarded by the two-sided filter)
- Symmetric, two-sided filter ( $v_{-l} = v_l$ ); filtering weights sum to one ( $\sum_{i=-m}^m v_i = 1$  with  $m = k = l$ )
- Constant and time trend in regression model invariant to filtering ( $X_F = X$ )

- Define filtered and unfiltered data sets as

$$\begin{aligned}y' &= (y_{-l}, \dots, y_{T+k}), \mathcal{T} \times 1, \\y_F &= Fy, (T+1) \times 1, \text{ and} \\y_U &= Uy, (T+1) \times 1; \\U &\equiv (\mathbf{O}_{(T+1) \times l} \mathbf{I}_{T+1} \mathbf{O}_{(T+1) \times k})\end{aligned}$$

- Matrix  $U$  cuts away observations at each end of the sample so that  $y_F$  and  $y_U$  are filtered and unfiltered data sets of equal sizes  $T+1$  with disturbances drawn from the same random vector  $\epsilon$

## Simple case: no lagged dependent variables

- Suppose  $y_t = x_t' \beta + \epsilon_t, \quad t = -l, \dots, 0, 1, \dots, T + k$
- $\Omega$  and  $F$  known, thus we know

$$\begin{aligned} E(\epsilon_U \epsilon_U') &\equiv \Omega_U = U \Omega U \\ E(\epsilon_F \epsilon_F') &\equiv \Omega_F = F \Omega F \end{aligned}$$

- $\hat{\beta}_j = (X_j' X_j)^{-1} X_j' y_j, \quad j = U, F$
- $\hat{\beta}$  unbiased:  $E(\hat{\beta}_F) = (X_F' X_F)^{-1} X_F' E(X_F \beta + F \epsilon) = \beta$
- and  $\text{Var}(\hat{\beta}_F) = (X_F' X_F)^{-1} X_F' \Omega_F X_F (X_F' X_F)^{-1}$

## Filtering ARIMA types of models

- Suppose

$$y_t = \mu + \gamma_t + z_t$$

and linear filter  $v(L)$  where  $v(L) = v(L^{-1})$  and  $v(1) = 1$

- Filtered version of  $y_t$  can be written as:

$$\begin{aligned} y_t &= z_t, & y_0 &= 0 \\ y_t^F &= v(L)z_t, & y_0^F &= 0 \end{aligned}$$

i.e.  $y_t^F = v(L)y_t$

- Then (consider simple case)  $y_t = \alpha y_{t-1} + \epsilon_t$
- $\rightarrow$  compare  $\hat{\alpha}$  with  $\hat{\alpha}_F$  to determine asymptotic bias

- Filtering of unit root processes:
  - when  $\phi(L)$  has root on unit circle,  $\hat{\alpha}$  and  $\hat{\alpha}_F$  both converge to one
  - thus, NO bias  $\rightarrow$  asymptotic null distributions of unit root test statistics not affected by filtering of the data
  - HOWEVER, finite sample distributions are affected by filtering
  
- Filtering Stationary ARMA processes:
  - when  $\phi(L)$  does not contain unit root  $\rightarrow$  asymptotic bias induced by filtering
  - Ghysels and Perron (1993) show: positive bias in case of linear X-11 filter (thus, unit root tests performed with filtered data can be expected to be less powerful against stationary alternatives)

# Conclusions

- Use of adjusted (filtered) data
  - discards important information about economic dynamics
  - possibly introduces severe biases
- Use of unfiltered data
  - incorporates misspecification risk
  - standard statistical inference procedures often result in close fits at seasonal frequencies but underemphasizes other frequencies (Sims, 1974, 1993); loss function might put a heavy penalty on misfitting the seasonal frequencies
  - Hansen and Sargent (1993) exploit further sources of misspecification; their findings seem to support Sims (1974, 1993)
  - however, debate still ongoing; depends on the context